Statistics 351 (Fall 2015) Prof. Michael Kozdron

Lecture #5: The Transformation Theorem

Reference. §1.2.1 pages 20–23

Last class we discussed how to perform a change-of-variables in the one-dimensional case. In multiple dimensions, the idea is the same, although things get notationally messier.

Recall. The one-dimensional change-of-variables formula is usually written as

$$\int_{a}^{b} f(h(x))h'(x) \, \mathrm{d}x = \int_{h(a)}^{h(b)} f(u) \, \mathrm{d}u,$$

i.e., u = h(x), du = h'(x) dx.

We can write this as

$$\int_{h(a)}^{h(b)} f(x) \, \mathrm{d}x = \int_{a}^{b} f(h(y)) h'(y) \, \mathrm{d}y. \tag{*}$$

It is in this form that there is a resemblance to the multidimensional formula.

Let $\mathbf{X} = (X_1, \ldots, X_n)'$ be a random vector, and let $g : \mathbb{R}^n \to \mathbb{R}^n$ be a sufficiently smooth function. That is, we write $g = (g_1, \ldots, g_n)$ with $g_i : \mathbb{R}^n \to \mathbb{R}$. Set $\mathbf{Y} = g(\mathbf{X})$ so that $(Y_1, \ldots, Y_n)' = (g_1(\mathbf{X}), \ldots, g_n(\mathbf{X}))'$ where $g_i(\mathbf{X}) = g_i(X_1, \ldots, X_n)$. Writing things out explicitly gives

$$Y_1 = g_1(X_1, \dots, X_n), \dots, Y_n = g_n(X_1, \dots, X_n).$$

Let $h = g^{-1}$ so that $\mathbf{X} = h(\mathbf{Y})$, and let

$$J = \left| \frac{\mathrm{d}(\overline{x})}{\mathrm{d}(\overline{y})} \right| = \begin{vmatrix} \frac{\partial x_1}{\partial y_1} & \frac{\partial x_1}{\partial y_2} & \cdots & \frac{\partial x_1}{\partial y_n} \\ \frac{\partial x_2}{\partial y_1} & \frac{\partial x_2}{\partial y_2} & \cdots & \frac{\partial x_2}{\partial y_n} \\ \vdots & \vdots & \ddots & \vdots \\ \frac{\partial x_n}{\partial y_1} & \frac{\partial x_n}{\partial y_2} & \cdots & \frac{\partial x_n}{\partial y_n} \end{vmatrix}$$

be the Jacobian.

Let $B \subset \mathbb{R}^n$. Our goal is to compute $P\{\mathbf{Y} \in B\}$ which we will do in two ways. The first way is as follows. Let $f_{\mathbf{Y}}(\overline{y})$ denote the density function for \mathbf{Y} so that

$$P\{\mathbf{Y} \in B\} = \int \cdots \int f_{\mathbf{Y}}(\overline{y}) \,\mathrm{d}\overline{y}. \tag{**}$$

The second way is by changing variables. Let $h = g^{-1}$ so that $h(B) = \{\overline{x} : g(\overline{x}) \in B\}$. Then, in analogy with the one-dimensional case, we have

$$P\{\mathbf{Y}\in B\} = P\{g(\mathbf{X})\in B\} = P\{\mathbf{X}\in g^{-1}(B)\} = P\{\mathbf{X}\in h(B)\} = \int_{h(B)} \cdots \int f_{\mathbf{X}}(\overline{x}) \,\mathrm{d}\overline{x}$$

and a change-of variables $\overline{y} = g(\overline{x})$ (i.e., $\overline{x} = h(\overline{y}), \ d\overline{x} = |J| d\overline{y}$) gives

$$\int_{h(B)} \cdots \int f_{\mathbf{X}}(\overline{x}) \, \mathrm{d}\overline{x} = \int_{B} \cdots \int_{B} f_{\mathbf{X}}(h(\overline{y})) \cdot |J| \, \mathrm{d}\overline{y} \qquad (***)$$

Note that if we are in one dimension, then (* * *) reduces to (*).

We now see that (*) and (**) give two distinct expressions for $P\{\mathbf{Y} \in B\}$, namely

$$P\{\mathbf{Y} \in B\} = \int \cdots \int f_{\mathbf{Y}}(\overline{y}) \, \mathrm{d}\overline{y} = \int \cdots \int f_{\mathbf{X}}(h(\overline{y})) \cdot |J| \, \mathrm{d}\overline{y}$$

Since this is true for any $B \subset \mathbb{R}^n$ we conclude that the integrands must be equal; that is,

$$f_{\mathbf{Y}}(\overline{y}) = f_{\mathbf{X}}(h(\overline{y})) \cdot |J|. \tag{(\dagger)}$$

In other words, this gives us a formula for the density of the random vector $\mathbf{Y} = g(\mathbf{X})$ in terms of the density of the random vector $\mathbf{X} = g^{-1}(\mathbf{Y}) = h(\mathbf{Y})$.

Example. Let $X, Y \in \mathcal{N}(0, 1)$ be independent. Prove that X+Y and X-Y are independent $\mathcal{N}(0, 2)$ random variables.

Solution. The fact that X + Y and X - Y each have a $\mathcal{N}(0, 2)$ distribution can be proved using moment generating functions as done in Stat 251. It is the *independence* of X + Y and X - Y that is newly proved.

Let U = X + Y and V = X - Y so that

$$X = \frac{U+V}{2} \quad \text{and} \quad Y = \frac{U-V}{2}.$$

That is, $\mathbf{X} = (X, Y)'$ and $\mathbf{Y} = (U, V)'$. We also find

- $g(\mathbf{X}) = g(X, Y) = (X + Y, X Y) = (U, V) = \mathbf{Y}$, and
- $h(\mathbf{Y}) = h(U, V) = \left(\frac{U+V}{2}, \frac{U-V}{2}\right) = (X, Y) = \mathbf{X}.$

We compute the Jacobian as

$$J = \begin{vmatrix} \frac{\partial x}{\partial u} & \frac{\partial x}{\partial v} \\ \frac{\partial y}{\partial u} & \frac{\partial y}{\partial v} \end{vmatrix} = \begin{vmatrix} 1/2 & 1/2 \\ 1/2 & -1/2 \end{vmatrix} = -1/2.$$

By the previous result (†), we conclude

$$\begin{aligned} f_{U,V}(u,v) &= f_{X,Y}(h(u,v)) \cdot |J| \\ &= f_{X,Y}((u+v)/2, (u-v)/2) \cdot 1/2 \\ &= 1/2 \cdot f_X((u+v)/2) \cdot f_Y((u-v)/2) & \text{by the assumed independence of } X \text{ and } Y \\ &= \frac{1}{2} \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}\frac{(u+v)^2}{4}} \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}\frac{(u-v)^2}{4}} \\ &= \frac{1}{\sqrt{2\pi \cdot 2}} e^{-\frac{1}{2}\frac{u^2}{2}} \cdot \frac{1}{\sqrt{2\pi \cdot 2}} e^{-\frac{1}{2}\frac{v^2}{2}} \end{aligned}$$

for $-\infty < u, v < \infty$. In other words, since $f_{U,V}(u, v) = f_U(u) \cdot f_V(v)$ we conclude that U = X + Y and V = X - Y are independent $\mathcal{N}(0, 2)$ random variables.