Statistics 351 Fall 2007
Independence of $\bar{X}$ and $S^{2}$
Theorem. Suppose that $X_{1}, \ldots, X_{n}$ are independent $\mathcal{N}(0,1)$ random variables. If

$$
\bar{X}=\frac{1}{n} \sum_{i=1}^{n} X_{i} \quad \text { and } \quad S^{2}=\frac{1}{n-1} \sum_{i=1}^{n}\left(X_{i}-\bar{X}\right)^{2}
$$

denote the sample mean and sample variance, respectively, then $\bar{X}$ and $S^{2}$ are independent.
Our proof of this theorem will follow the outline proposed by Exercise V.8.2.
Proof. Since $X_{1}, \ldots, X_{n}$ are i.i.d. $\mathcal{N}(0,1)$, we conclude (using, say moment generating functions) that $\bar{X} \in \mathcal{N}(0,1 / n)$. Similarly, we can show that

$$
\begin{aligned}
\bar{X}-X_{j}=\frac{1}{n}\left(X_{1}+\cdots+X_{j-1}+X_{j+1}+\cdots+X_{n}\right)-\frac{n-1}{n} X_{j} & \in \mathcal{N}\left(0, \frac{n-1}{n^{2}}+\frac{(n-1)^{2}}{n^{2}}\right) \\
& =\mathcal{N}\left(0, \frac{n-1}{n}\right)
\end{aligned}
$$

and so

$$
X_{j}-\bar{X} \in \mathcal{N}\left(0, \frac{n-1}{n}\right)
$$

as well. We also note that

$$
\operatorname{Cov}\left(X_{j}, \bar{X}\right)=\operatorname{Cov}\left(X_{j}, \frac{1}{n} \sum_{i=1}^{n} X_{i}\right)=\frac{1}{n} \sum_{i=1}^{n} \operatorname{Cov}\left(X_{j}, X_{i}\right)=\frac{1}{n} \operatorname{Cov}\left(X_{j}, X_{j}\right)=\frac{1}{n},
$$

and so for $j \neq k$ it follows that

$$
\begin{aligned}
\operatorname{Cov}\left(X_{j}-\bar{X}, X_{k}-\bar{X}\right) & =\operatorname{Cov}\left(X_{j}, X_{k}\right)-\operatorname{Cov}\left(X_{j}, \bar{X}\right)-\operatorname{Cov}\left(\bar{X}, X_{k}\right)+\operatorname{Cov}(\bar{X}, \bar{X}) \\
& =0-\frac{1}{n}-\frac{1}{n}+\frac{1}{n} \\
& =-\frac{1}{n}
\end{aligned}
$$

using the fact that $\operatorname{Cov}(\bar{X}, \bar{X})=\operatorname{Var}(\bar{X})=1 / n$. Similarly,

$$
\begin{aligned}
\operatorname{Cov}\left(\bar{X}, X_{j}-\bar{X}\right) & =\operatorname{Cov}\left(X_{j}, \bar{X}\right)-\operatorname{Cov}(\bar{X}, \bar{X}) \\
& =\frac{1}{n}-\frac{n}{n^{2}} \\
& =0
\end{aligned}
$$

Thus, we see that $\left(\bar{X}, X_{1}-\bar{X}, \ldots, X_{n}-\bar{X}\right)^{\prime} \in \mathcal{N}(\overline{0}, \boldsymbol{\Lambda})$ where

$$
\boldsymbol{\Lambda}=\left[\begin{array}{ccccc}
1 / n & 0 & 0 & \cdots & 0 \\
0 & (n-1) / n & -1 / n & \cdots & -1 / n \\
0 & -1 / n & (n-1) / n & \cdots & -1 / n \\
\vdots & \vdots & \vdots & \ddots & \vdots \\
0 & -1 / n & -1 / n & \cdots & (n-1) / n
\end{array}\right]
$$

By a theorem from last lecture (Theorem V.7.2), we conclude from the form of $\Lambda$ that $\bar{X}$ and $\left(X_{1}-\bar{X}, \ldots, X_{n}-\bar{X}\right)^{\prime}$ are independent normal random vectors. It now follows from the transformation theorem (Theorem I.2.1 on page 23) that since $\mathbf{X}=\left(X_{1}-\bar{X}, \ldots, X_{n}-\bar{X}\right)^{\prime}$ and $\bar{X}$ are independent, so too are

$$
\bar{X} \text { and } \mathbf{X}^{\prime} \mathbf{X}=\sum_{i=1}^{n}\left(X_{i}-\bar{X}\right)^{2}
$$

This now implies that $\bar{X}$ and $S^{2}$ are independent.

