Stat 252.01 Winter 2007 Assignment #8 Solutions

1. As always,  $\alpha = P_{H_0}(\text{reject } H_0)$  and  $\beta = P_{H_A}(\text{accept } H_0)$ . Since X has an Exponential $(\lambda)$  distribution so that  $f(x|\lambda) = \lambda^{-1}e^{-x/\lambda}$ , x > 0, and since our rejection region is  $\{X < c\}$ , we find that

$$\alpha = P_{H_0}(\text{reject } H_0) = P(X < c | \lambda = 1) = \int_0^c e^{-x} dx = 1 - e^{-c}$$

and

$$\beta = P_{H_A}(\text{accept } H_0) = P(X > c | \lambda = 1/2) = \int_c^\infty 2e^{-2x} \, dx = e^{-2c}.$$

Thus,  $\alpha = 1 - e^{-c}$  and  $\beta = e^{-2c}$  which easily implies that

$$1 - \alpha = \sqrt{\beta}.$$

Rewrite this as  $\alpha + \sqrt{\beta} = 1$  to illustrate the direct tradeoff between them: as  $\alpha$  increases,  $\beta$  must decrease, and vice-versa.

**2.** If X has an Exponential( $\lambda$ ) distribution so that  $f(x|\lambda) = \lambda^{-1}e^{-x/\lambda}$ , x > 0, then the Fisher information is

$$I(\lambda) = \frac{1}{\lambda^2}$$

and the maximum likelihood estimator is

$$\hat{\lambda}_{\text{MLE}} = \overline{X}.$$

Hence, a significance level 0.1 test of  $H_0: \lambda = 1/5$  vs.  $H_A: \lambda \neq 1/5$  has rejection region

$$RR = \left\{ \left| \sqrt{nI(\hat{\lambda}_{\text{MLE}})} \left( \hat{\lambda}_{\text{MLE}} - \lambda_0 \right) \right| \ge z_{0.05} \right\}$$

or

$$RR = \left\{ \frac{\sqrt{n}}{\overline{X}} \left| \overline{X} - \frac{1}{5} \right| \ge 1.645 \right\}.$$

(10.10) Let  $\mu$  denote the average hardness index. In order to test the manufacturer's claim, we want to test  $H_0: \mu \ge 64$  against  $H_A: \mu < 64$ . It is equivalent to test  $H_0: \mu = 64$  against  $H_A: \mu < 64$ . The test statistic is given by

$$Z = \frac{\overline{Y} - \mu_0}{\sigma/\sqrt{n}} = \frac{62 - 64}{8/\sqrt{50}} \approx -1.77.$$

In order to conduct this test at the significance level  $\alpha = 0.01$ , we find that the rejection region is

$$RR = \{Z < z_{0.01} = -2.326\}.$$

Since Z = -1.77 does not fall in the rejection region (-1.77 > -2.326), we do not reject  $H_0$  in favour of  $H_A$  at the 0.01 level. Thus, we conclude that there is insufficient evidence to reject the manufacturer's claim.

(10.38) The rejection region is

$$\frac{\hat{\theta} - \theta_0}{\sigma_{\hat{\theta}}} < -z_\alpha$$

which is true if and only if

$$\ddot{\theta} + z_{\alpha}\sigma_{\hat{\theta}} < \theta_0.$$

That is,  $H_0$  will be rejected at the significance level  $\alpha$  if and only if the  $100(1 - \alpha)\%$  upper confidence bound for  $\theta$  (namely,  $\hat{\theta} + z_{\alpha}\sigma_{\hat{\theta}}$ ) is less than  $\theta_0$ .

(10.50) A *t*-test can be used whenever one wants to conduct a hypothesis test of the population mean when the population is known to have a normal distribution with unknown variance. The *t*-test also works reasonably well for populations whose distribution is mound-shaped (and resembles the normal).

(10.73) Let  $\sigma$  denote the standard deviation of the accuracy of the precision instrument. In order to assess the precision, we want to test  $H_0: \sigma = 0.7$  against  $H_A: \sigma > 0.7$ . It is equivalent to test  $H_0: \sigma^2 = 0.49$  against  $H_A: \sigma^2 > 0.49$ . The sample variance is given by

$$s^{2} = \frac{1}{n-1} \sum_{i=1}^{n} (y_{i} - \overline{y})^{2} = \frac{(353 - 352.5)^{2} + (351 - 352.5)^{2} + (351 - 352.5)^{2} + (355 - 352.5)^{2}}{3} \approx 3.667$$

so that the test statistic is

$$\chi^2 = \frac{(n-1)s^2}{\sigma_0^2} \approx \frac{3 \cdot 3.667}{0.49} \approx 22.45.$$

(Recall that this is the distribution of the null hypothesis, i.e., assuming that  $\sigma_0^2 = 0.49$ .) Consulting Table 6 for the  $\chi^2$  density, we find  $\chi^2_{0.005,3} = 12.8381$ . Since  $22.45 \gg 12.8381$ , we see that the *p*-value must be smaller than 0.005.