

Math/Stat 251 Fall 2015

Some Examples of a One-Dimensional Change of Variables (November 2, 2015)

Suppose that X is a continuous random variable and that $Y = g(X)$ for some continuous function $g : \mathbb{R} \rightarrow \mathbb{R}$ so that Y is itself a continuous random variable. It is often the case in practice that one knows the density function of X and seeks the density function of Y . Fortunately, if g is a nice function (as it usually is in practice), then it is straightforward to determine the density of Y from first principles. Basically, one starts with the definition of the distribution function of Y substitutes in $Y = g(X)$, and solves for X . This produces an integral expression involving the density function of X which can then be differentiated using the fundamental theorem of calculus to yield the density function for Y . Sometimes this is called a *one-dimensional change of variables*. The following examples illustrate this technique. Remember that in order to use the fundamental theorem of calculus, it must be the case that a variable appears in the upper limit of integration and that no variable appears in the lower limit of integration.

Example. Suppose that $X \sim \mathcal{N}(0, 1)$. Let $Y = e^X$. Determine the density function of Y .

Solution. Let $Y = e^X$. For $y > 0$, the distribution function of Y is

$$F_Y(y) = \mathbf{P}(Y \leq y) = \mathbf{P}(e^X \leq y) = \mathbf{P}(X \leq \log y) = \int_{-\infty}^{\log y} \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx$$

so that the density function of Y is

$$f_Y(y) = \frac{d}{dy} F_Y(y) = \frac{d}{dy} \int_{-\infty}^{\log y} \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx = \frac{1}{\sqrt{2\pi}} e^{-(\log y)^2/2} \cdot \frac{d}{dy} \log y = \frac{1}{y\sqrt{2\pi}} e^{-(\log y)^2/2}$$

for $y > 0$. *The random variable Y is an example of a log-normal random variable which is regularly encountered in the mathematical theory of stock option pricing.*

Example. Suppose that $X \sim \mathcal{N}(0, 1)$. Let $Y = X^2$. Determine the density function of Y .

Solution. Let $Y = X^2$ so that

$$F_Y(y) = \mathbf{P}(Y \leq y) = \mathbf{P}(X^2 \leq y).$$

Note that since X can take on *any* real value, we have

$$\begin{aligned} \mathbf{P}(X^2 \leq y) &= \mathbf{P}(-\sqrt{y} \leq X \leq \sqrt{y}) = \int_{-\sqrt{y}}^{\sqrt{y}} f_X(x) dx \\ &= \int_{-\sqrt{y}}^0 f_X(x) dx + \int_0^{\sqrt{y}} f_X(x) dx \\ &= \int_0^{\sqrt{y}} f_X(x) dx - \int_0^{-\sqrt{y}} f_X(x) dx \\ &= \int_0^{\sqrt{y}} \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx - \int_0^{-\sqrt{y}} \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx, \end{aligned}$$

and so

$$\begin{aligned}
 f_Y(y) &= \frac{d}{dy} F_Y(y) = \frac{d}{dy} \int_0^{\sqrt{y}} \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx - \frac{d}{dy} \int_0^{-\sqrt{y}} \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx \\
 &= \frac{1}{\sqrt{2\pi}} e^{-(\sqrt{y})^2/2} \cdot \frac{d}{dy} (\sqrt{y}) - \frac{1}{\sqrt{2\pi}} e^{-(-\sqrt{y})^2/2} \cdot \frac{d}{dy} (-\sqrt{y}) \\
 &= \frac{1}{\sqrt{2\pi}} e^{-y/2} \cdot \frac{1}{2\sqrt{y}} + \frac{1}{\sqrt{2\pi}} e^{-y/2} \cdot \frac{1}{2\sqrt{y}} \\
 &= \frac{1}{\sqrt{2\pi}} \frac{1}{\sqrt{y}} e^{-y/2}
 \end{aligned}$$

for $y > 0$. Note that the random variable Y has a $\text{Gamma}(1/2, 1/2)$ distribution, or equivalently, $Y \sim \chi^2(1)$ and often appears in statistical inference. It also sometimes appears in the mathematical theory of stock option pricing.

Example. Suppose that $X \in \Gamma(a, b)$ so that the density of X is

$$f_X(x) = \frac{b^a}{\Gamma(a)} x^{a-1} e^{-bx}$$

for $x \geq 0$. Let $Y = 1/X$. Determine the density function of Y .

Solution. Let $Y = 1/X$. For $y > 0$, the distribution function of Y is

$$\begin{aligned}
 F_Y(y) &= \mathbf{P}(Y \leq y) = \mathbf{P}(1/X \leq y) = \mathbf{P}(X \geq 1/y) = 1 - \mathbf{P}(X < 1/y) \\
 &= 1 - \int_{-\infty}^{1/y} \frac{b^a}{\Gamma(a)} x^{a-1} e^{-bx} dx
 \end{aligned}$$

so that the density function of Y is

$$f_Y(y) = \frac{d}{dy} F_Y(y) = \frac{d}{dy} \left(1 - \int_{-\infty}^{1/y} \frac{b^a}{\Gamma(a)} x^{a-1} e^{-bx} dx \right) = \frac{b^a}{\Gamma(a)} y^{1-a} e^{-b/y} \cdot \frac{1}{y^2} = \frac{b^a}{\Gamma(a)} y^{-a-1} e^{-b/y}$$

for $y > 0$. The random variable Y is an example of an inverse gamma random variable with parameters a and b and is used primarily in Bayesian statistics though it sometimes finds applications in actuarial science.